

Lecture 4: 4D-Var Diagnostics

Outline

- Posterior/analysis error estimates
- Consistency checks, hypothesis tests, degrees of freedom & information content
- Array modes
- Clipped analyses

Posterior Error Estimates

Posterior/Analysis Error Estimates

Posterior error covariance:

$$\begin{aligned} E^a &= \langle (z_a - z_t)(z_a - z_t)^T \rangle \\ &= \langle (z_b + \delta z_a - z_t)(z_b + \delta z_a - z_t)^T \rangle \\ &= (\mathbf{I} - \mathbf{K}\mathbf{G})\mathbf{D}(\mathbf{I} - \mathbf{K}\mathbf{G})^T + \mathbf{K}\mathbf{R}\mathbf{K}^T \end{aligned}$$

(Caveat: E^a correct *only if* \mathbf{D} and \mathbf{R} are correct)

But for dual Lanczos vector formulation:

$$\mathbf{K} \approx \tilde{\mathbf{K}}_k = \mathbf{D}\mathbf{G}^T\mathbf{R}^{-1/2}\mathbf{V}_k\mathbf{T}_k^{-1}\mathbf{V}_k^T\mathbf{R}^{-1/2}$$

Lanczos vectors are orthonormal, and normal to subspace neglected by $\tilde{\mathbf{K}}_k$:

$$E^a \approx \tilde{E}^a = (\mathbf{I} - \tilde{\mathbf{K}}\mathbf{G})\mathbf{D}$$

Posterior/Analysis Error Estimates

Approx. posterior error covariance:

$$\begin{aligned} E^a &\approx \tilde{E}^a = (\mathbf{I} - \tilde{\mathbf{K}}\mathbf{G})\mathbf{D} \\ \tilde{E}^a &= \underbrace{(\mathbf{I} - \mathbf{D}\mathbf{G}^T\mathbf{R}^{-1/2}\mathbf{V}_k\mathbf{T}_k^{-1}\mathbf{V}_k^T\mathbf{R}^{-1/2}\mathbf{G})\mathbf{D}} \end{aligned}$$

Everything is available during inner-loops of 4D-PSAS and R4D-Var, at no extra cost

$$\tilde{E}^a \sim N_{\text{model}} \times N_{\text{model}} \quad \text{HUGE!}$$

Diagonal elements – *posterior* variances

define `POSTERIOR_ERROR_I`

Cross-covariance information from EOFs of \tilde{E}^a

define `POSTERIOR_EOFS`

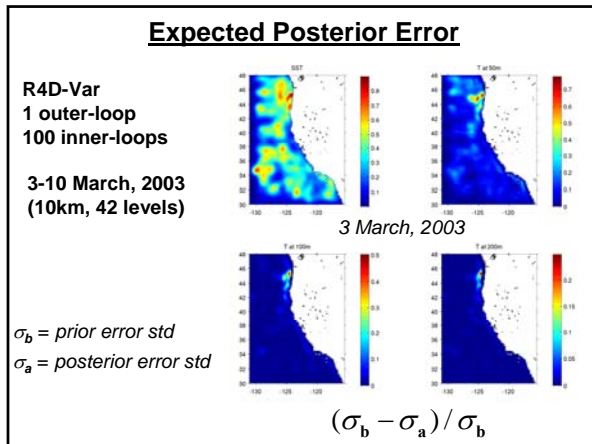
Posterior/Analysis Error Estimates

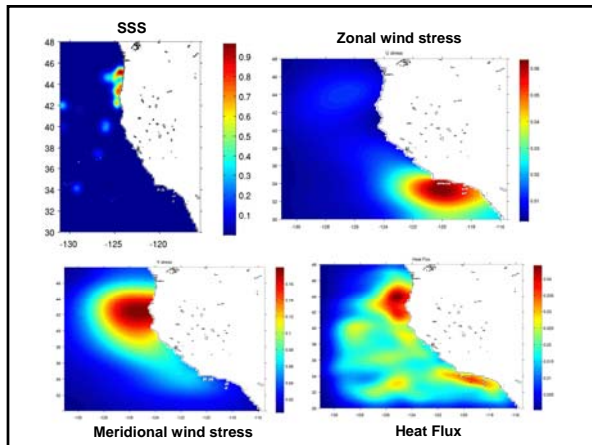
For the primal formulation:

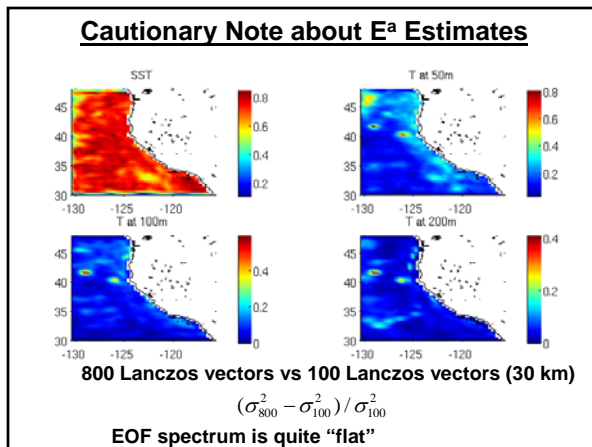
$$E^a = (\mathbf{D}^{-1} + \mathbf{G}^T\mathbf{R}^{-1}\mathbf{G})^{-1}$$

$$E^a \approx \tilde{E}^a = \mathbf{D}^{1/2}\mathbf{V}_k\mathbf{T}_k^{-1}\mathbf{V}_k^T\mathbf{D}^{1/2}$$

Straightforward but not yet implemented in ROMS I4D-Var (due to large I/O requirements)

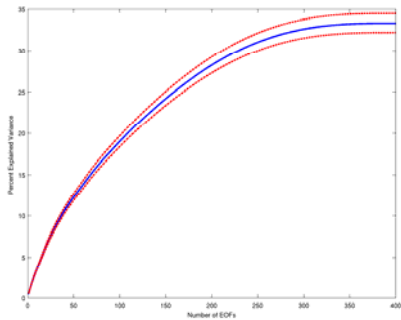






Expected Posterior Error EOFs

Flat spectrum - perhaps due to preconditioning which clusters eigenvalues of $(R^{1/2}GDG^T R^{1/2} + I)$ around 1.



Consistency checks,
hypothesis tests, degrees of
freedom & information content

Consistency Checks in Obs Space

Statistics of the innovation vectors following
Desroziers et al (2005):

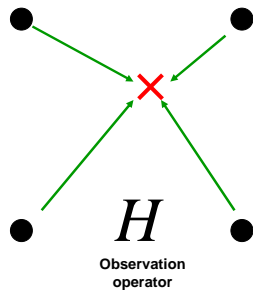
$$\mathbf{d} = (\mathbf{y} - H(\mathbf{x}_b))$$

$$\mathbf{d}_a^o = (\mathbf{y} - H(\mathbf{x}_a))$$

$$\mathbf{d}_b^a = (H(\mathbf{x}_a) - H(\mathbf{x}_b))$$

$$\tilde{\sigma}_b^2 = (\mathbf{d}_b^a)^T \mathbf{d} / p$$

$$\tilde{\sigma}_o^2 = (\mathbf{d}_a^o)^T \mathbf{d} / p$$



Compare $\tilde{\sigma}_o$ with σ_o & $\tilde{\sigma}_b$ with σ_b

Consistency Checks in Obs Space

Statistics of the innovation vectors following

Desroziers et al (2005):

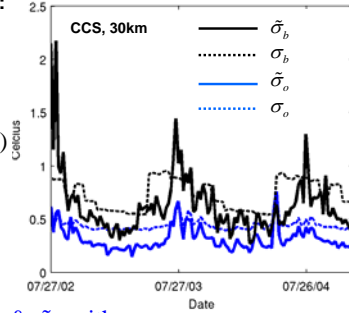
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$$\tilde{\sigma}_o^2 = (\mathbf{d}_a^o)^T \mathbf{d} / p$$



Compare $\tilde{\sigma}_o$ with σ_o & $\tilde{\sigma}_b$ with σ_b

Hypothesis Tests & Degrees of Freedom

Recall that the optimal increments minimize:

$$J = \underbrace{\frac{1}{2} \delta \mathbf{z}^T \mathbf{D}^{-1} \delta \mathbf{z}}_{J_b} + \underbrace{\frac{1}{2} (\mathbf{G} \delta \mathbf{z} - \mathbf{d})^T \mathbf{R}^{-1} (\mathbf{G} \delta \mathbf{z} - \mathbf{d})}_{J_o}$$

Theoretical min: $J_{\min} = N_{\text{obs}} / 2$

No. of dof in obs $\rightarrow (J_b)_{\min} = \text{Tr}(\mathbf{K}\mathbf{G}) / 2$

No. of dof in prior $\rightarrow (J_o)_{\min} = (N_{\text{obs}} - \text{Tr}(\mathbf{K}\mathbf{G})) / 2$

“dof” – degrees of freedom

(Bennett et al, 1993;
Cardinali et al, 2004;
Desroziers et al., 2009)

Degrees of Freedom & Information Content

Degrees of freedom in the obs:

$$(J_b)_{\min} = \text{Tr}(\mathbf{K}\mathbf{G}) / 2$$

But for $m = N_{\text{obs}}$:

$$\mathbf{K} = \mathbf{D}\mathbf{G}^T \mathbf{R}^{-1/2} \mathbf{V}_m \mathbf{T}_m^{-1} \mathbf{V}_m^T \mathbf{R}^{-1/2}$$

So:

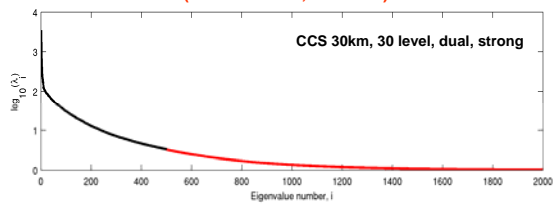
$$(J_b)_{\min} = \frac{1}{2} (N_{\text{obs}} - \text{Tr}(\mathbf{T}_m^{-1}))$$

$$= \frac{1}{2} \left(N_{\text{obs}} - \sum_{i=1}^{N_{\text{obs}}} \lambda_i^{-1} \right)$$

Eigenvalues
of \mathbf{T}_m^{-1}

Degrees of Freedom & Information Content

(LhessianEV=T, s4dvar.in)

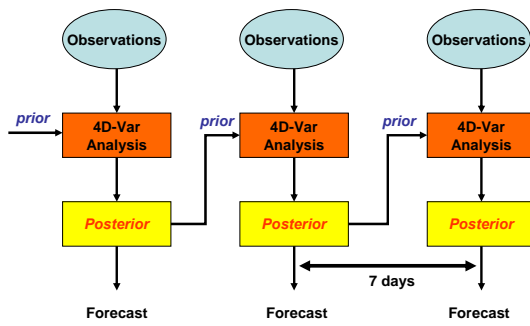


— Computed directly during R4D-Var

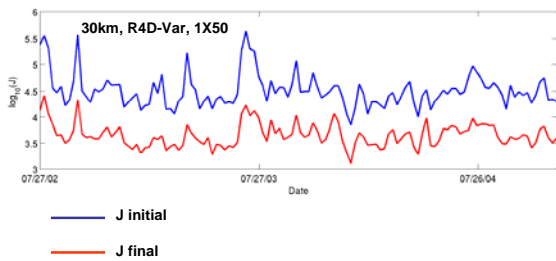
— Computed from a curve fit

$$\log_{10}(\lambda_i) = ae^{-bi}, a = 1.684, b = 2.7 \times 10^{-3}$$

Sequential 4D-Var with 30km CCS ROMS



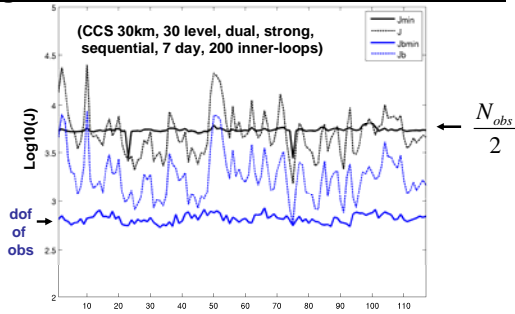
Sequential 4D-Var CCS ROMS



— J initial

— J final

Degrees of Freedom & Information Content



- Less than 10% of all observations provide independent info
- LOTS OF REDUNDANCY!
- $J_j > (J_j)_{min}$ and indicates over fitting to the obs
- $J \neq J_{min}$ and indicates that *prior* hypotheses are incorrect

Array Modes

The eigenvalues λ_i of T_m are also the eigenvalues of $(R^{1/2}GDG^T R^{1/2} + I)$, the preconditioned "stabilized representer matrix."

Consider the eigenpairs: (λ_i, \hat{w}_i)

Following Bennett (1985):

$$\delta x_a(t) = \mathcal{R}(t) R^{-1/2} W \Lambda^{-1} W^T R^{-1/2} d \quad \text{Analysis state-vector increment}$$

$$= \sum_{j=1}^{N_{obs}} \lambda_j^{-1} (\hat{w}_j^T R^{-1/2} d) \Psi_j(t)$$

$$\Psi_j(t) = \sum_{j=1}^{N_{obs}} w'_{ji} r_j(t) \quad \text{"array modes"}$$

$$w'_i = (w'_{ji}) = R^{-1/2} \hat{w}_i$$

Array Modes

State-vector analysis increment:

$$\delta x_a(t) = \sum_{i=1}^{N_{obs}} \lambda_i^{-1} (\hat{w}_i^T R^{-1/2} d) \Psi_i(t)$$

Largest eigenvalue λ_1 associated with $\Psi_1(t)$, nominally contributes least to analysis increment.

Thus $\Psi_1(t)$ represents the most stable component of δx_a with respect to changes in d .

Array modes can be readily computed after dual 4D-Var:

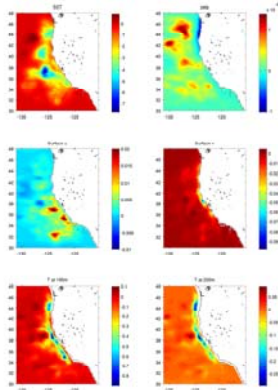
$$\Psi_j(t) = \mathcal{M}_b(t, t_0) DG^T R^{-1/2} \hat{w}_j$$

(define ARRAY_MODES)

Example:

CCS, 10km
R4D-Var
3-7 March, 2003

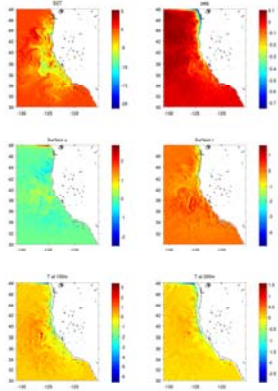
Ψ_1 3 March



Example:

CCS, 10km
R4D-Var
3-7 March, 2003

Ψ_1 7 March

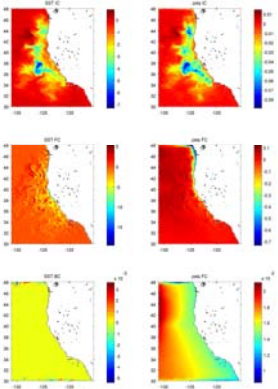


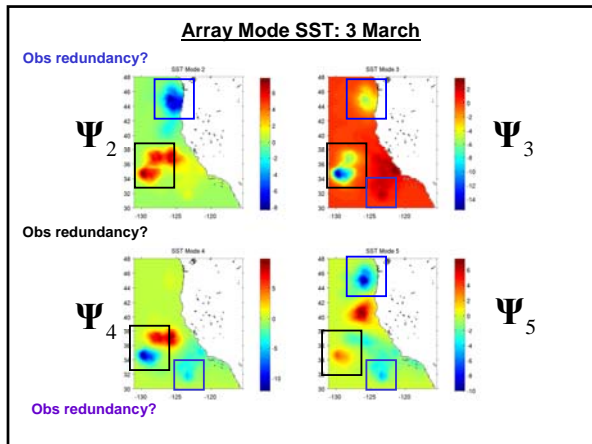
Example:

CCS, 10km
R4D-Var
3-7 March, 2003

Ψ_1 3 March

Control vector
contributions





Clipped Analyses

Following Bennett (1985)

$$\delta \mathbf{x}_a(t) = \sum_{i=1}^M \lambda_i^{-1} (\hat{\mathbf{w}}_i^T \mathbf{R}^{-1/2} \mathbf{d}) \Psi_i(t)$$

Truncate ("clip") the summation, discarding array modes Ψ_i for $i > M$

M is based on a criteria that reflects information content (i.e. $\lambda_i < \alpha \lambda_1$ for $i > M$).

α chosen so Ψ_i have scales $<$ model resolution for $i > M$

Clipping using Indirect Representer Algorithm

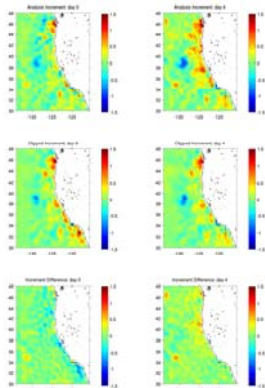
Clipped analyses can be computed AFTER running R4D-Var or 4D-PSAS by exploiting the indirect representer algorithm according to:

$$\delta \mathbf{x}_a(t) = \mathcal{M}_b(t, t_0) \mathbf{D} \mathbf{G}^T \sum_{i=1}^M \lambda_i^{-1} (\hat{\mathbf{w}}_i^T \mathbf{R}^{-1/2} \mathbf{d}) \hat{\mathbf{w}}_i$$

where $\mathcal{M}_b(t, t_0)$ is TLROMS linearized about the prior/background.

(Moore et al, 2010b)

Example:
 CCS, 30km
 R4D-Var
 3-7 March, 2003
 1 outer-loop
 800 inner-loops



Clipped using
 "1% rule" of
 Bennett & McIntosh
 (1985)

Discard Ψ_i for which

$$\lambda_i < 0.01\lambda_1$$

$$M = 255$$

Issues, Things to do, & Coming Soon

- *Posterior* error estimates available for 4D-PSAS & R4D-Var - available soon for I4D-Var
- *Posterior* error estimates are likely over estimates due to span of observation or model space by Lanczos vectors
- (4D-Var)^T provides more reliable estimates (Lecture 5)

Summary

- *Posterior* error estimates available via the utility afforded by the Lanczos algorithm.
- *Posterior* errors tend to be overestimates since $m \ll N_{obs}$
- Consistency checks and hypothesis tests are useful indicators of validity of *prior* hypotheses and info content.
- Array modes provide info about possible data redundancy.
- Clipping can remove unphysical features from analyses.

References

- [Bennett, A.F., 1985: Array design by inverse methods. *Prog. Oceanogr.*, **15**, 129-156.](#)
- [Bennett, A.F. and P.C. McIntosh, 1982: Open ocean modelling as an inverse problem: tidal theory. *J. Phys. Oceanogr.*, **12**, 1004-1018.](#)
- [Bennett, A.F., L. Leslie, C. Hagelberg and P. Powers, 1993: Tropical cyclone prediction using a barotropic model initialized by a generalized inverse method. *Mon. Wea. Rev.*, **121**, 1714-1729.](#)
- [Cardinali, C., S. Pezzulli and E. Andersson, 2004: Influence-matrix diagnostic of a data assimilation system. *Q. J. R. Meteorol. Soc.*, **130**, 2767-2786.](#)
- [Desroziers, G., L. Berre, B. Chapnik and P. Poli, 2005: Diagnosis of observation, background and analysis-error statistics in observation space. *Q. J. R. Meteorol. Soc.*, **131**, 3385-3396.](#)
- [Desroziers, G., L. Berre, V. Chabot and B. Chapnik, 2009: A posteriori diagnostics in an ensemble of perturbed analyses. *Mon. Wea. Rev.*, **137**, 3420-3436.](#)

References

- [Moore, A.M., H.G. Arango, G. Broquet, B.S. Powell, J. Zavala-Garay and A.T. Weaver, 2010a: The Regional Ocean Modeling System \(ROMS\) 4-dimensional data assimilation systems: Part I – System overview. *Ocean Modelling*, Submitted.](#)
- [Moore, A.M., H.G. Arango, G. Broquet, C. Edwards, M. Veneziani, B.S. Powell, D. Foley, J. Doyle, D. Costa and P. Robinson, 2010b: The Regional Ocean Modeling System \(ROMS\) 4-dimensional data assimilation systems: Part II – Performance and application to the California Current System. *Ocean Modelling*, Submitted.](#)
